

The background of the slide features a large, faint watermark of the University of Washington seal. The seal is circular and contains the text 'UNIVERSITY OF WASHINGTON' around the perimeter and '1861' at the bottom. In the center of the seal is a shield with a star above it and a crest with a book and a torch. The shield is flanked by two figures holding a banner. The text 'UW' is displayed in a dark blue box in the top-left corner.

**UW**

# **Simulation-Based Estimation of Continuous Time Models in R**

R/Finance 2010

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Joint work with:

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University of Hawaii

# Introduction

Goal: Estimate parameters of continuous time diffusion model from discretely sampled data

$$dy_t = \alpha(y_t, \theta)dt + \sigma(y_t, \theta)dW_t, \quad dW_t \sim \text{iid } N(0, dt)$$

Examples

$$\text{OU: } dy_t = (\theta_0 - \theta_1 y_t)dt + \theta_2 dW_t, \quad \alpha(y_t, \theta) = \theta_0 - \theta_1 y_t, \quad \sigma(y_t, \theta) = \theta_2$$

$$\text{CIR: } dy_t = (\theta_0 - \theta_1 y_t)dt + \theta_2 \sqrt{y_t} dW_t, \quad \alpha(y_t, \theta) = \theta_0 - \theta_1 y_t, \quad \sigma(y_t, \theta) = \theta_2 \sqrt{y_t}$$

$$\text{GCIR: } dy_t = (\theta_0 - \theta_1 y_t)dt + \theta_2 y_t^\gamma dW_t,$$

$$\alpha(y_t, \theta) = \theta_0 - \theta_1 y_t, \quad \sigma(y_t, \theta) = \theta_2 y_t^\gamma$$

# Estimation Methods

- MLE – often not feasible
- MLE of approximated model – difficult
- QMLE of discretized model – easy but biased
- GMM – inefficient and biased
- Bayesian MCMC Methods - promising
- Indirect Inference – Corrects bias in QMLE
  - focus of talk

# Indirect Inference

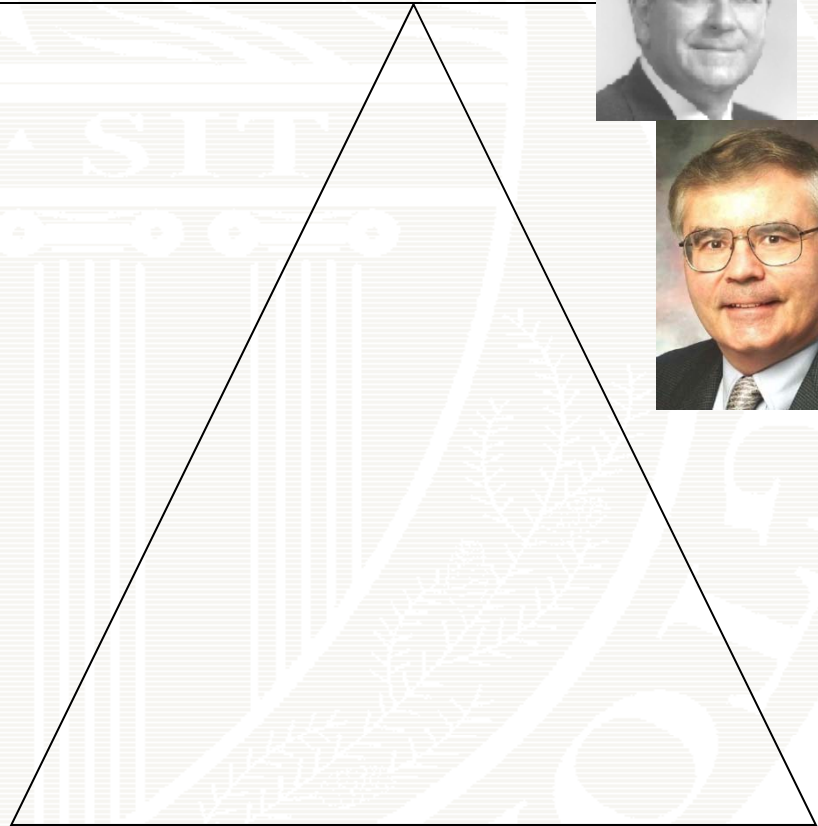
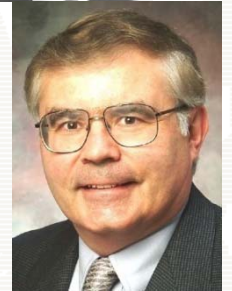
- Distance-based methodology (aka II) developed by Smith (1993), Gouriéroux, Monfort, and Renault (1993)
- Score-based methodology (aka EMM) developed by Gallant and Tauchen (1996)



II



EMM



1861



II



EMM



• **Computationally less intensive**  
(Gallant and Tauchen, 1996;  
Chumacero, 2001)

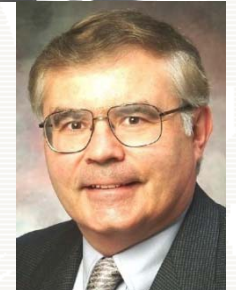




II



EMM



- **Smaller bias and MSE**  
in MA models (Ghysels, Khalaf and Vodounou, 2003)

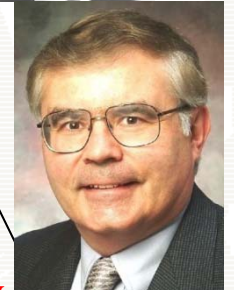
- **Computationally less intensive**  
(Gallant and Tauchen, 1996, Chumacero, 2001)



||

- **Smaller bias and MSE**  
in MA models (Ghysels, Khalaf and Vodounou, 2003)
- **More accurate inference**  
for AR models  
(Duffee and Stanton, 2008)

EMM



- **Computationally less intensive**  
(Gallant and Tauchen, 1996;  
Chumacero, 2001)



# Research Agenda and R Contribution

- Implement indirect inference estimation techniques for some commonly used continuous time models (e.g., OU, CIR, etc.)
- Provide systematic comparison and evaluation of different estimators
- Create indirectInference R package
- Give practical advice on use of techniques

# Indirect Inference Set-up

$\{y_t\}_{t=\Delta}^{n\Delta}$  observations with observation interval  $\Delta$

Structural model:  $F_\theta$ ,  $\theta \in \mathbb{R}^p$ , stationary and ergodic

Auxiliary model:  $\tilde{F}_\mu$ ,  $\mu \in \mathbb{R}^r$ ,  $r \geq p$

$\tilde{\mu} = \arg \max_{\mu} Q_n(\{y_t\}_{t=\Delta}^{n\Delta}, \mu)$ , where

$$\tilde{Q}_n = \frac{1}{n - m} \sum_{t=(m+1)\Delta}^{n\Delta} \tilde{f}(y_t; x_{t-\Delta}, \mu), \quad x_{t-\Delta} = \{y_i\}_{i=t-m\Delta}^{t-\Delta}$$

$\tilde{f}(y_t; x_{t-\Delta}, \mu) =$  conditional log density of  $y_t$  for the model  $\tilde{F}_\mu$

$\mu(\theta) = \arg \max_{\mu} E_{F_\theta}[\tilde{f}(y_t; x_{t-\Delta}, \mu)] = p \lim \tilde{\mu}$  under  $F_\theta$   
 = binding function

# Example: OU Model

$$F_\theta : dy_t = (\theta_0 - \theta_1)y_t dt + \theta_2 dW_t, \quad \theta_i > 0, p=3, \Delta=1/52$$

$$y_t = \frac{\theta_0}{\theta_1} (1 - e^{-\theta_1 \Delta}) + e^{-\theta_1 \Delta} y_{t-\Delta} + \theta_2 \sqrt{\frac{1 - e^{-2\theta_1 \Delta}}{2\theta_1}} z_t, \quad z_t \sim \text{iid } N(0,1)$$

$$\tilde{F}_\mu : y_t = y_{t-\Delta} + (\mu_0 + \mu_1 y_{t-\Delta}) \Delta + \mu_2 \sqrt{\Delta} \varepsilon_{t-\Delta}$$

$$= \mu_0 \Delta + (1 - \mu_2 \Delta) y_{t-\Delta} + \mu_2 \sqrt{\Delta} \varepsilon_{t-\Delta}, \quad \varepsilon_{t-\Delta} \sim \text{iid } N(0,1), r=3$$

$$\mu(\theta) = p \lim \tilde{\mu} \neq \theta$$

$$\mu_0(\theta) = \frac{\theta_0}{\theta_1 \Delta} (1 - e^{-\theta_1 \Delta}), \quad \mu_1(\theta) = \frac{1}{\Delta} (1 - e^{-\theta_1 \Delta}), \quad \mu_2(\theta) = \theta_2 \sqrt{\frac{1 - e^{-2\theta_1 \Delta}}{2\theta_1 \Delta}}$$

# Example: OU Model

- Estimating the “crude Euler” auxiliary model leads to biased estimates (Lo, 1988)
  - Asymptotic discretization bias =  $\mu(\theta) - \theta$
  - $\mu(\theta) - \theta \rightarrow 0$  as  $\Delta \rightarrow 0$
- $\mu(\theta)$  is invertible giving analytic II estimates

$$\hat{\theta}^{II} = \mu^{-1}(\tilde{\mu})$$

$$\hat{\theta}_0^{II} = \frac{\tilde{\mu}_0}{\tilde{\mu}_1 \Delta} \ln(1 - \tilde{\mu}_1 \Delta), \quad \hat{\theta}_1^{II} = \frac{-1}{\Delta} \ln(1 - \tilde{\mu}_1 \Delta),$$

$$\hat{\theta}_2^{II} = \tilde{\mu}_2 \sqrt{\frac{2 \ln(1 - \tilde{\mu}_1 \Delta)}{1 - e^{\ln(1 - \tilde{\mu}_1 \Delta)}}}$$

# Non-simulation based Estimation

- Assume  $\mu(\theta)$  is known (very rare!)
- EMM is GMM with population moment

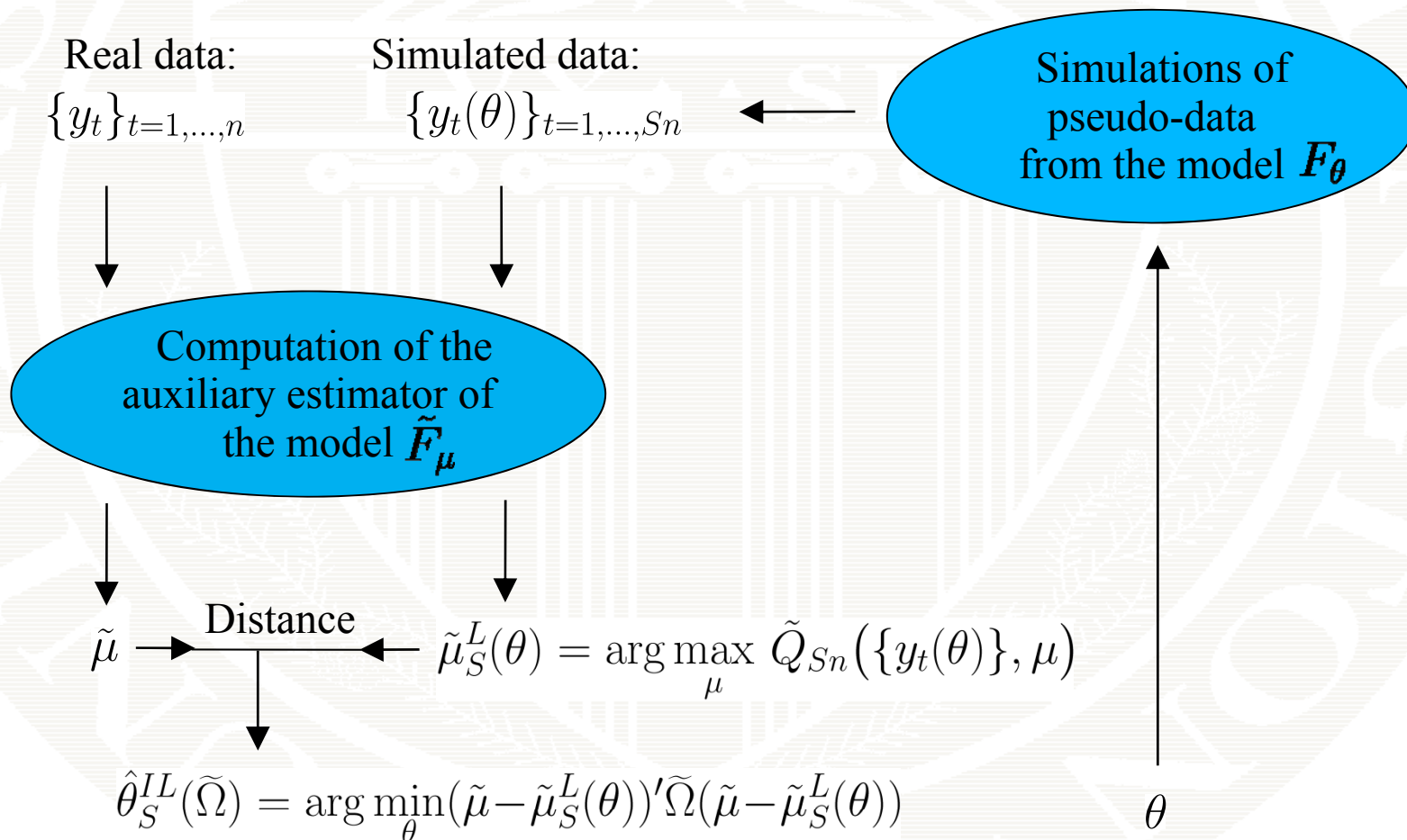
$$E_{F_\theta} \left[ \frac{\partial \tilde{f}(y_t; x_{t-\Delta}, \mu)}{\partial \mu} \right]_{\mu=\mu(\theta)} = 0$$

- It minimizes distance between  $\mu(\theta)$  and  $\tilde{\mu}$
- Asymptotically equivalent to MLE when auxiliary model encompasses structural model

# Simulation-based EMM and II

- $\mu(\theta)$  is unknown
- $\tilde{\mu}$  is used to estimate  $\mu(\theta_{true})$
- With EMM, simulations for a given  $\theta$  are used to approximate the expectation of sample score
- With II, simulations are used to approximate  $\mu(\theta)$  for any  $\theta$
- Gouriéroux and Monfort (1996) describe 3 types of II estimators and 2 types of EMM estimators

# 1<sup>st</sup> Type of Simulation-based II: IL



# 2<sup>nd</sup> Type of II: IM

Real data:

$$\{y_t\}_{t=1, \dots, n}$$

Simulated data:

$$\{y_t^s(\theta)\}_{t=1, \dots, n}^{s=1, \dots, S}$$

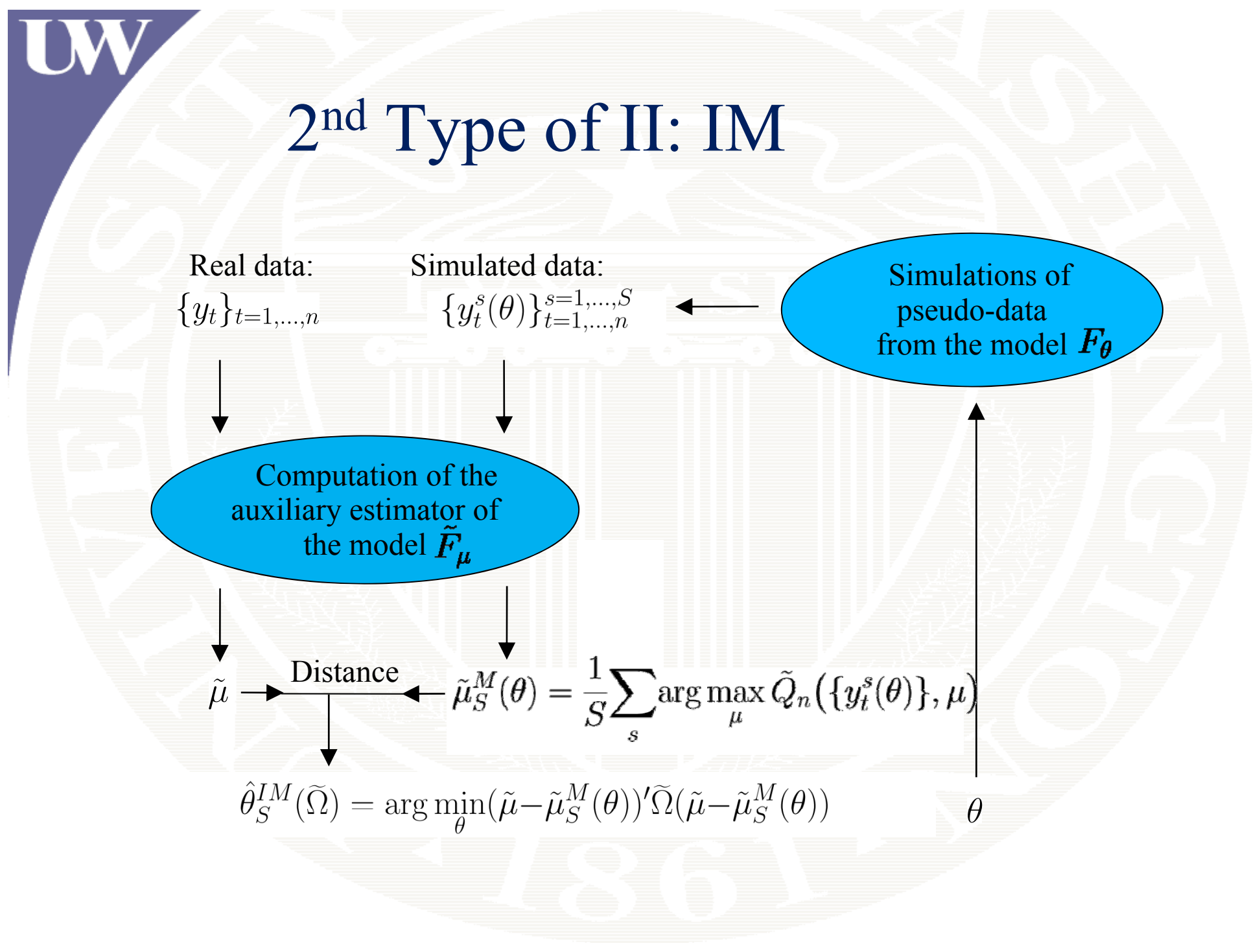
Simulations of pseudo-data from the model  $F_\theta$

Computation of the auxiliary estimator of the model  $\tilde{F}_\mu$

Distance

$$\tilde{\mu} \quad \longleftrightarrow \quad \tilde{\mu}_S^M(\theta) = \frac{1}{S} \sum_s \arg \max_\mu \tilde{Q}_n(\{y_t^s(\theta)\}, \mu)$$

$$\hat{\theta}_S^{IM}(\tilde{\Omega}) = \arg \min_\theta (\tilde{\mu} - \tilde{\mu}_S^M(\theta))' \tilde{\Omega} (\tilde{\mu} - \tilde{\mu}_S^M(\theta))$$





# 3<sup>rd</sup> Type of II: IA

Real data:

$$\{y_t\}_{t=1, \dots, n}$$

Simulated data:

$$\{y_t^s(\theta)\}_{t=1, \dots, n}^{s=1, \dots, S}$$

Simulations of pseudo-data from the model  $F_\theta$

Computation of the auxiliary estimator of the model  $\tilde{F}_\mu$

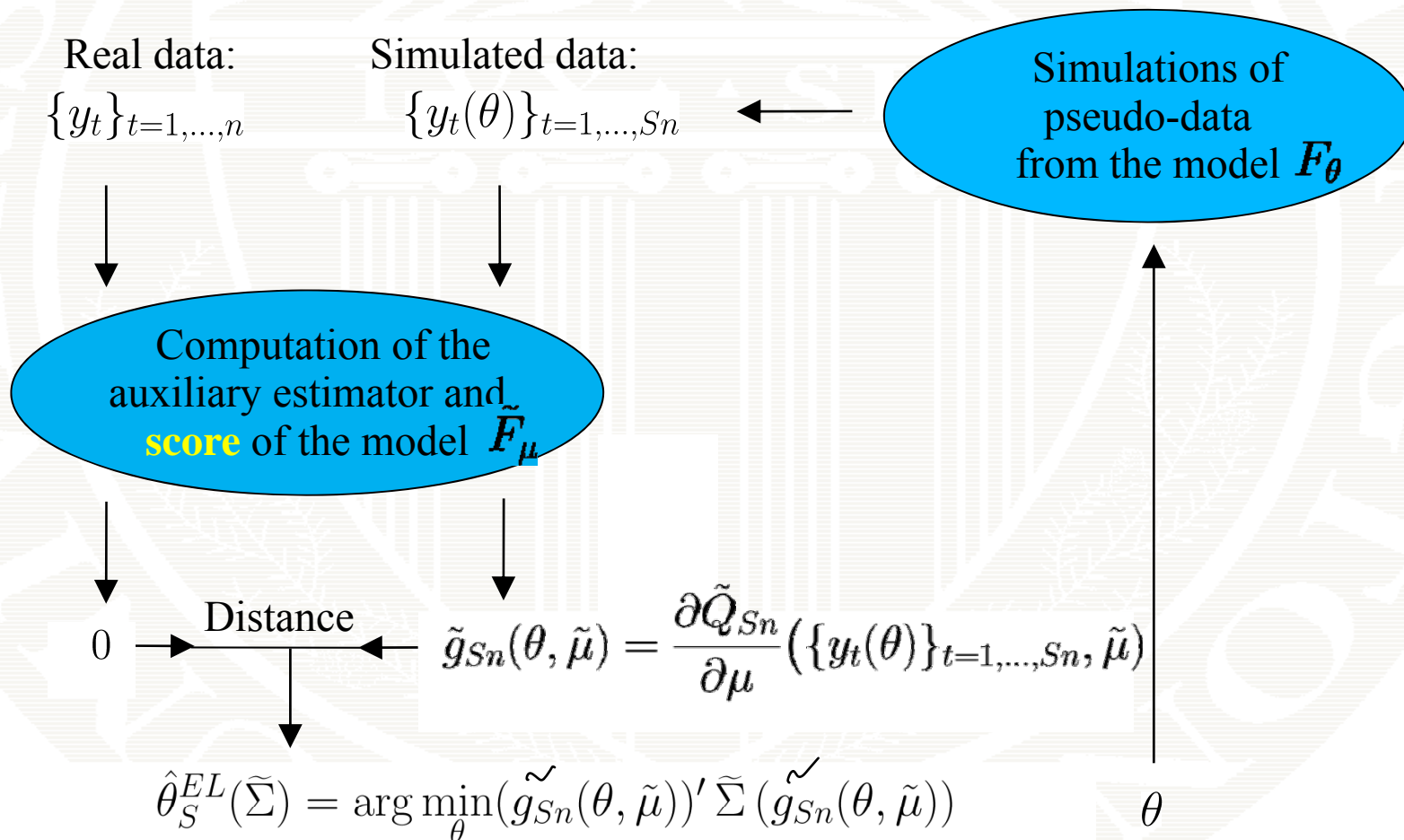
Distance

$$\tilde{\mu} \quad \longleftrightarrow \quad \tilde{\mu}_S^A(\theta) = \arg \max_{\mu} \frac{1}{S} \sum_s \tilde{Q}_n(\{y_t^s(\theta)\}, \mu)$$

$$\hat{\theta}_S^{IA}(\tilde{\Omega}) = \arg \min_{\theta} (\tilde{\mu} - \tilde{\mu}_S^A(\theta))' \tilde{\Omega} (\tilde{\mu} - \tilde{\mu}_S^A(\theta))$$

$\theta$

# 1<sup>st</sup> Type of EMM: EL



# 2<sup>nd</sup> Type of EMM: EA

Real data:  
 $\{y_t\}_{t=1,\dots,n}$

Simulated data:  
 $\{y_t^s(\theta)\}_{t=1,\dots,n}^{s=1,\dots,S}$

Simulations of pseudo-data from the model  $F_\theta$

Computation of the auxiliary estimator and **score** of the mode  $\tilde{F}_\mu$

Distance

$$g_n^s(\theta, \tilde{\mu}) = \frac{\partial \tilde{Q}_n}{\partial \mu}(\{y_t^s(\theta)\}_{t=1,\dots,n}, \tilde{\mu})$$

$$\hat{\theta}_S^{EA}(\tilde{\Sigma}) = \arg \min_{\theta} \left( \frac{1}{S} \sum_{s=1}^S g_n^s(\theta, \tilde{\mu}) \right)' \tilde{\Sigma} \left( \frac{1}{S} \sum_{s=1}^S g_n^s(\theta, \tilde{\mu}) \right)$$

# R Implementation of II

- Estimate Euler auxiliary model parameters  $\mu$  from observed data  $\{y_t\}$  by QMLE

$$y_{t+\Delta} - y_t = \alpha(y_t, \mu)\Delta + \sigma(y_t, \mu)\sqrt{\Delta}z_t, \quad z_t \sim \text{iid } N(0,1)$$

$$\tilde{\mu} = \arg \max_{\mu} Q_n \left( \{y_t\}_{t=\Delta}^{n\Delta}, \mu \right), \text{ where}$$

$$\tilde{Q}_n = \frac{1}{n-m} \sum_{t=(m+1)\Delta}^{n\Delta} \tilde{f}(y_t; x_{t-\Delta}, \mu), \quad x_{t-\Delta} = \{y_i\}_{i=t-m\Delta}^{t-\Delta}$$

- Use function `EULERloglik()` from R package `sde`
- Use R function `optim()`

# R Implementation of II

- Simulate from  $F_\theta$ 
  - In general, cannot do exact simulations because transition density is not known
  - Simulate from very fine Euler discretization
  - Use function `sde.sim()` from R package `sde`
  - Use custom C code for fast simulation
  - Need to worry about “inadmissible” or “explosive” simulations from inappropriate  $\theta$  – need to “bullet proof” the simulator

## R Implementation of II

- For distance-based II, estimate binding function  $\mu(\theta)$  from simulated data  $\{y_t^s(\theta)\}$

$$\tilde{\mu}_S^L = \arg \max_{\mu} Q_{Sn} \left( \{y_t^s(\theta)\}, \mu \right), \text{ where}$$

$$\tilde{Q}_{Sn} = \frac{1}{n'} \sum_{t=(m+1)\Delta}^{Sn\Delta} \tilde{f}(y_t^s(\theta); x_{t-\Delta}^s(\theta), \mu),$$

- Use same random number seed for all  $\theta$

## R Implementation of II

- For score-based II, estimate auxiliary score from simulated data  $\{y_t^s(\theta)\}$  and evaluate at auxiliary parameter estimate

$$g_{Sn}(\{y_t^s(\theta)\}, \tilde{\mu}) = \frac{\partial Q_{Sn}(\{y_t^s(\theta)\}, \tilde{\mu})}{\partial \mu}$$

- User specified function to evaluate score function
- Use same random number seed for all  $\theta$

## R Implementation of II

- For distance-based II, estimate  $\theta$

$$\hat{\theta}_S^{II} = \arg \min_{\theta} (\tilde{\mu} - \tilde{\mu}_S^i(\theta))' \tilde{\Omega} (\tilde{\mu} - \tilde{\mu}_S^i(\theta)), \quad i = L, M$$

- For score-based II, estimate

$$\hat{\theta}_S^{EMM} = \arg \min_{\theta} g_{Sn}(\theta, \tilde{\mu})' \tilde{\Sigma} g_{Sn}(\theta, \tilde{\mu})'$$

- If  $p = r$  then use identity matrix for weight matrix
- For optimization, use R function `optim()` with Nelder-Meade simplex algorithm



# Illustration

- OU Process calibrated to US interest rates used by Phillips and Yu (2009)

$$\theta_0 = 0.01, \theta_1 = 0.10, \theta_2 = 0.10$$

$$\frac{\theta_0}{\theta_1} = 0.10 = \text{annualized avg rate,}$$

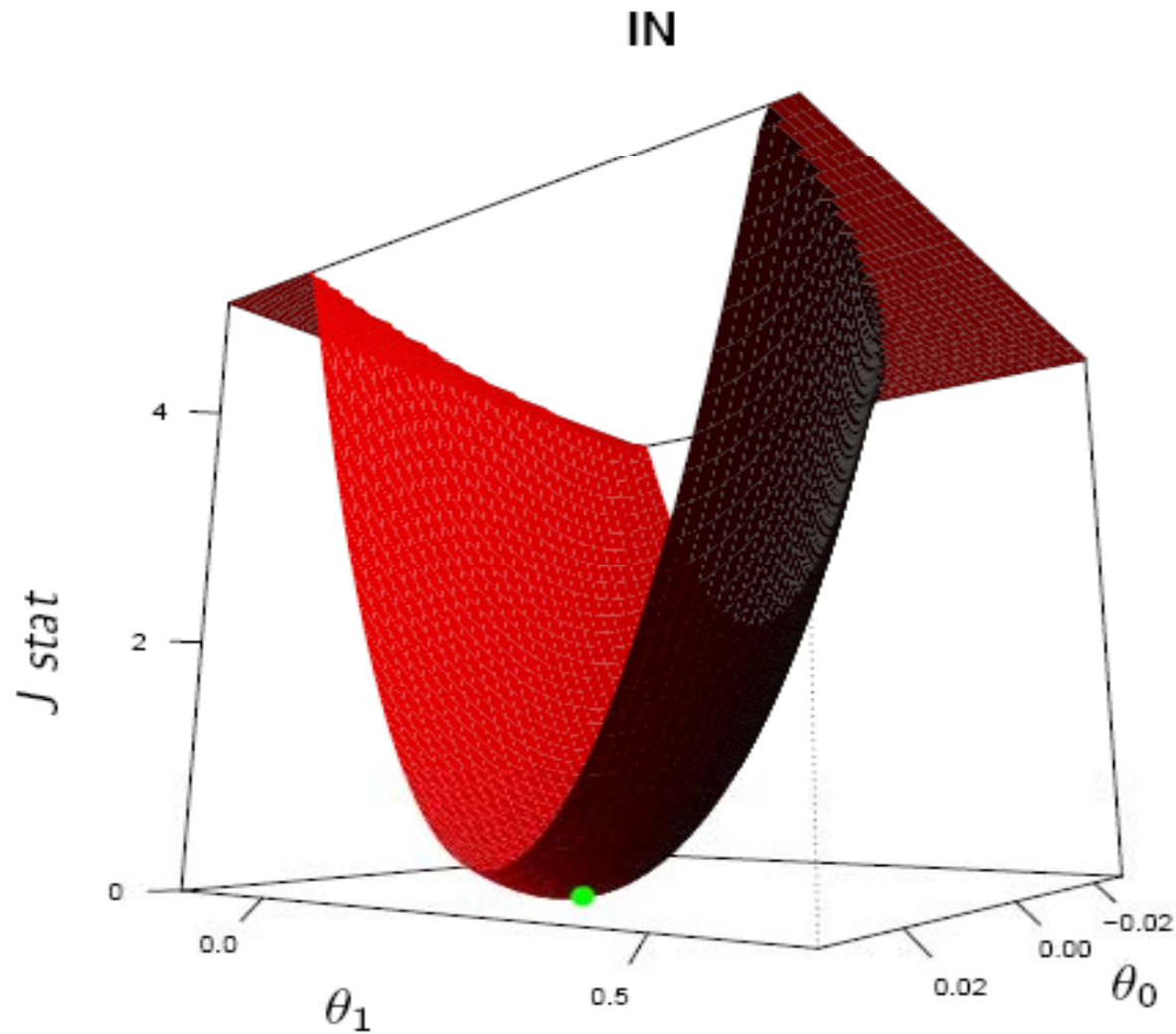
$$\theta_1 = 0.1 \Rightarrow 7 \text{ year half of rate shock}$$

$$\theta_2 = 0.10 = \text{annualized rate volatility}$$

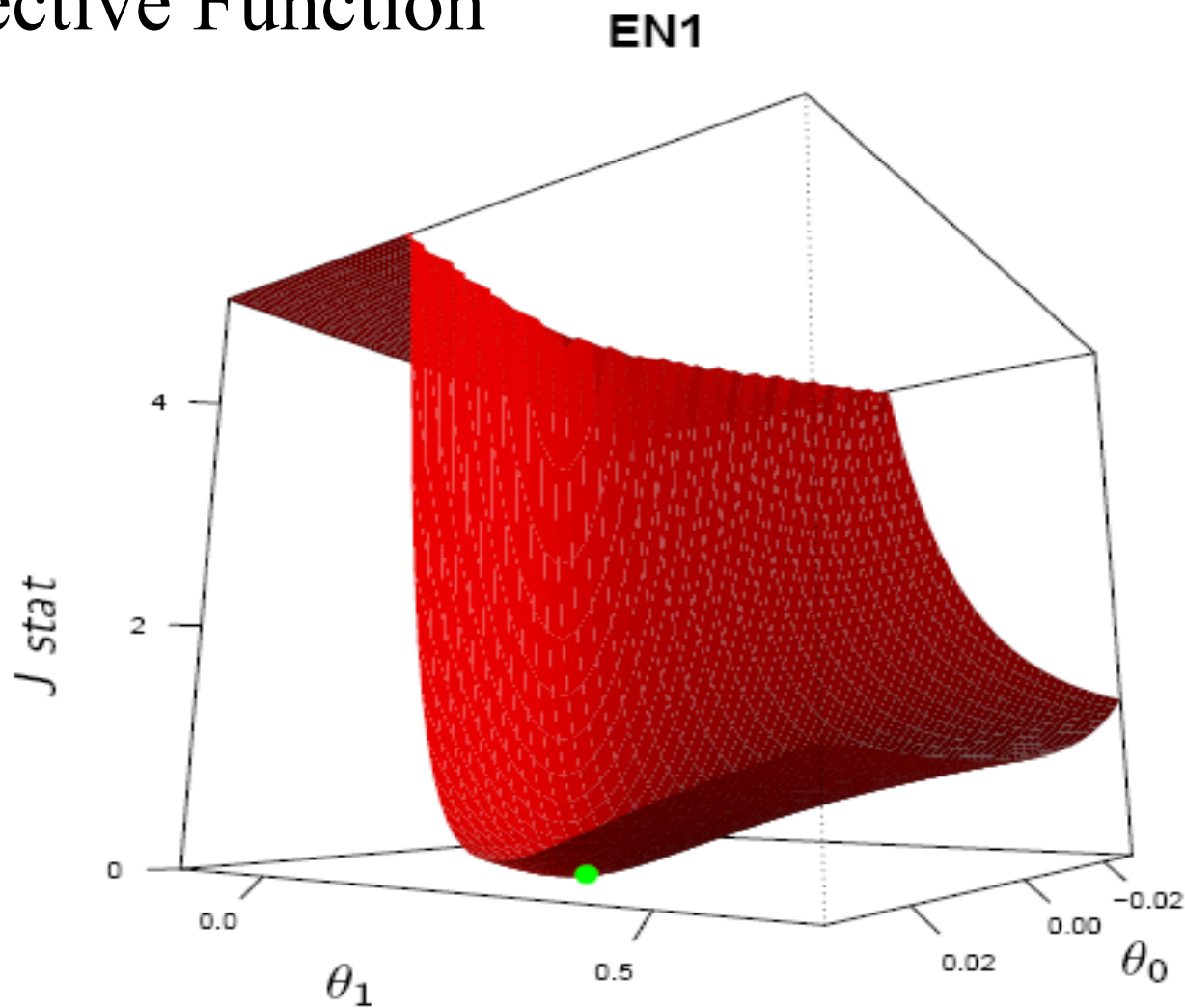
$$T = 19.23, \Delta = 1/52 \Rightarrow n = 1000$$

- $\theta_1$  is the most difficult parameter to estimate

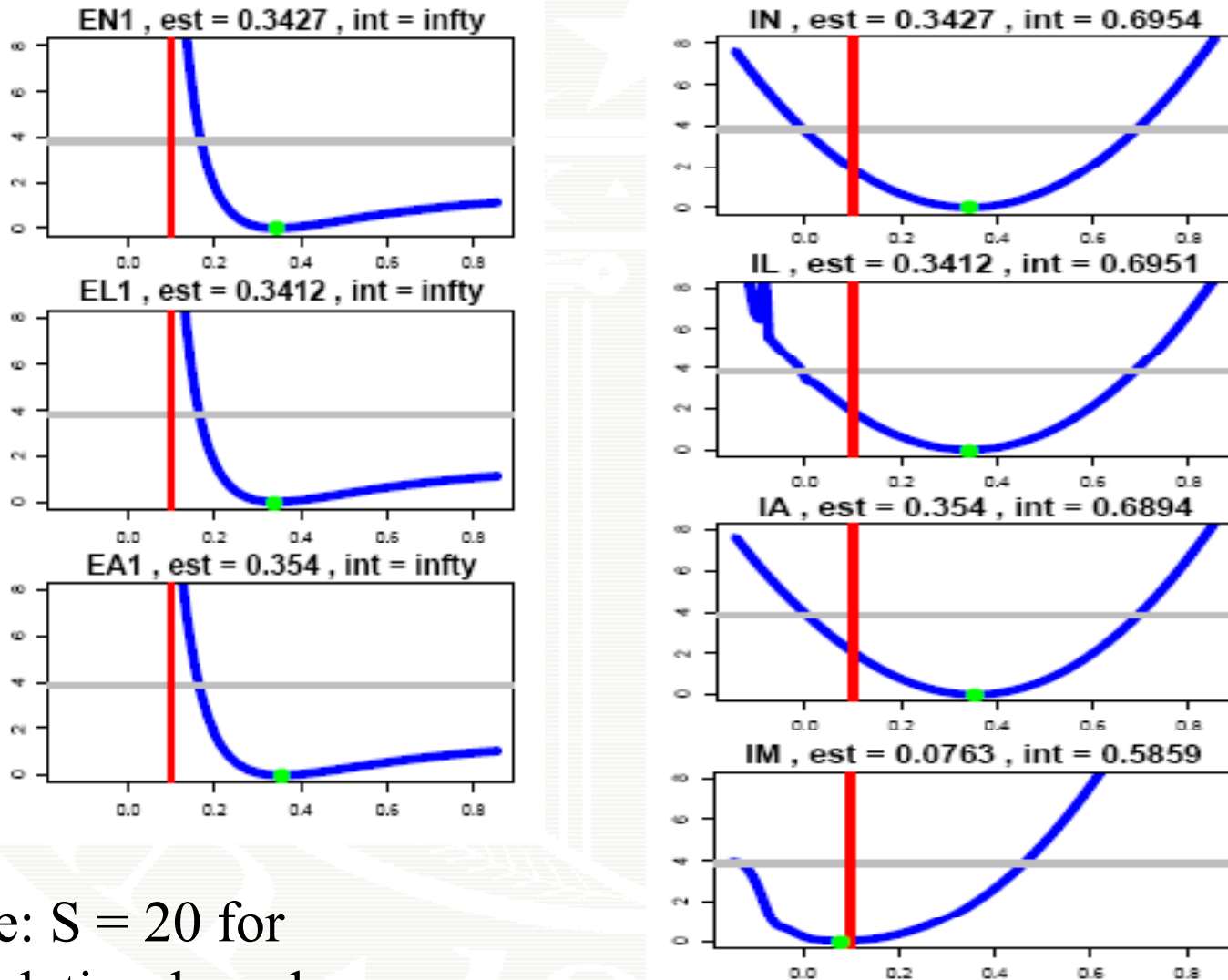
# Shape of distance-based II Objective function



# Shape of Gallant-Tauchen Score-based II Objective Function

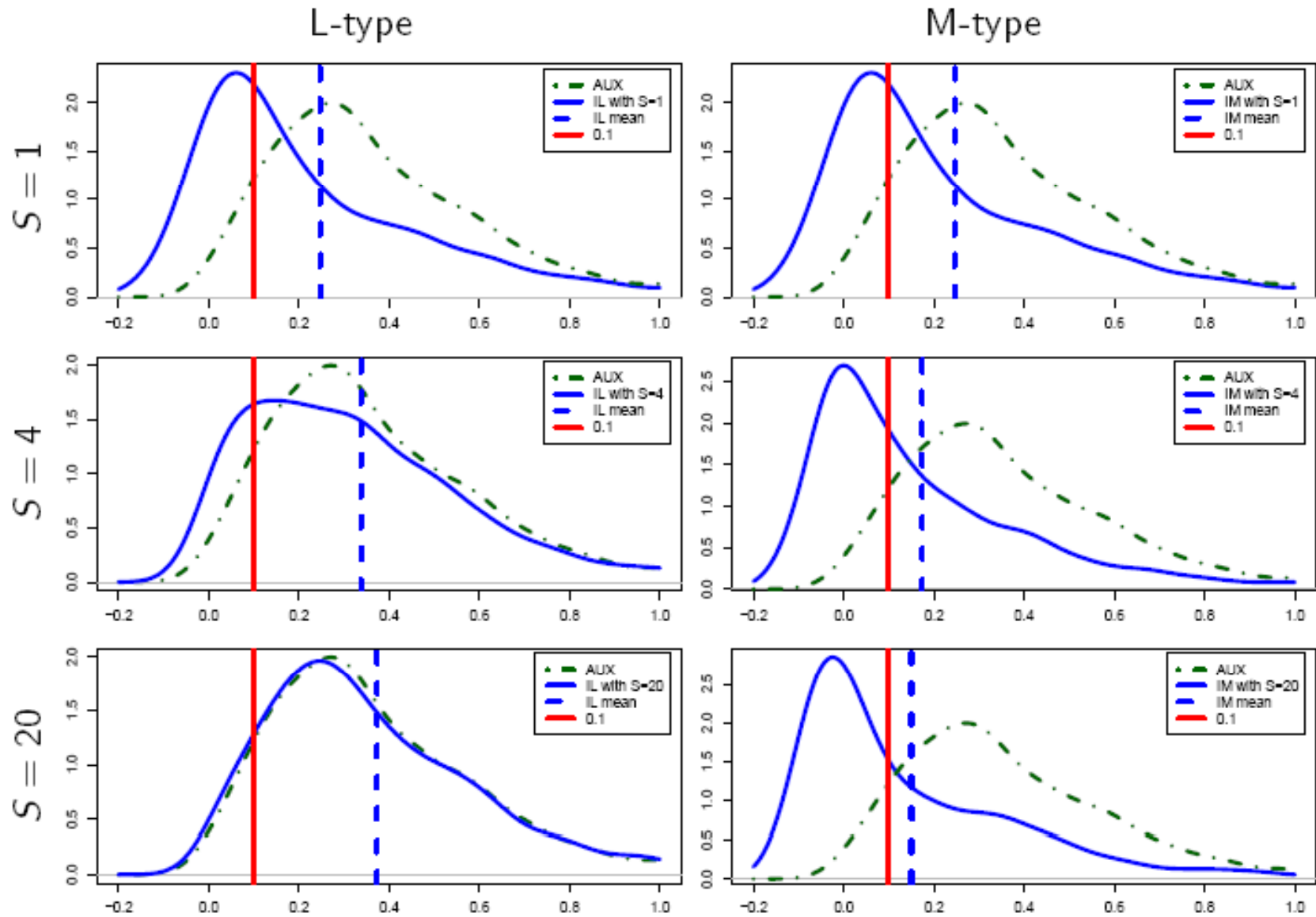


# 95% Confidence Intervals for $\theta_1$

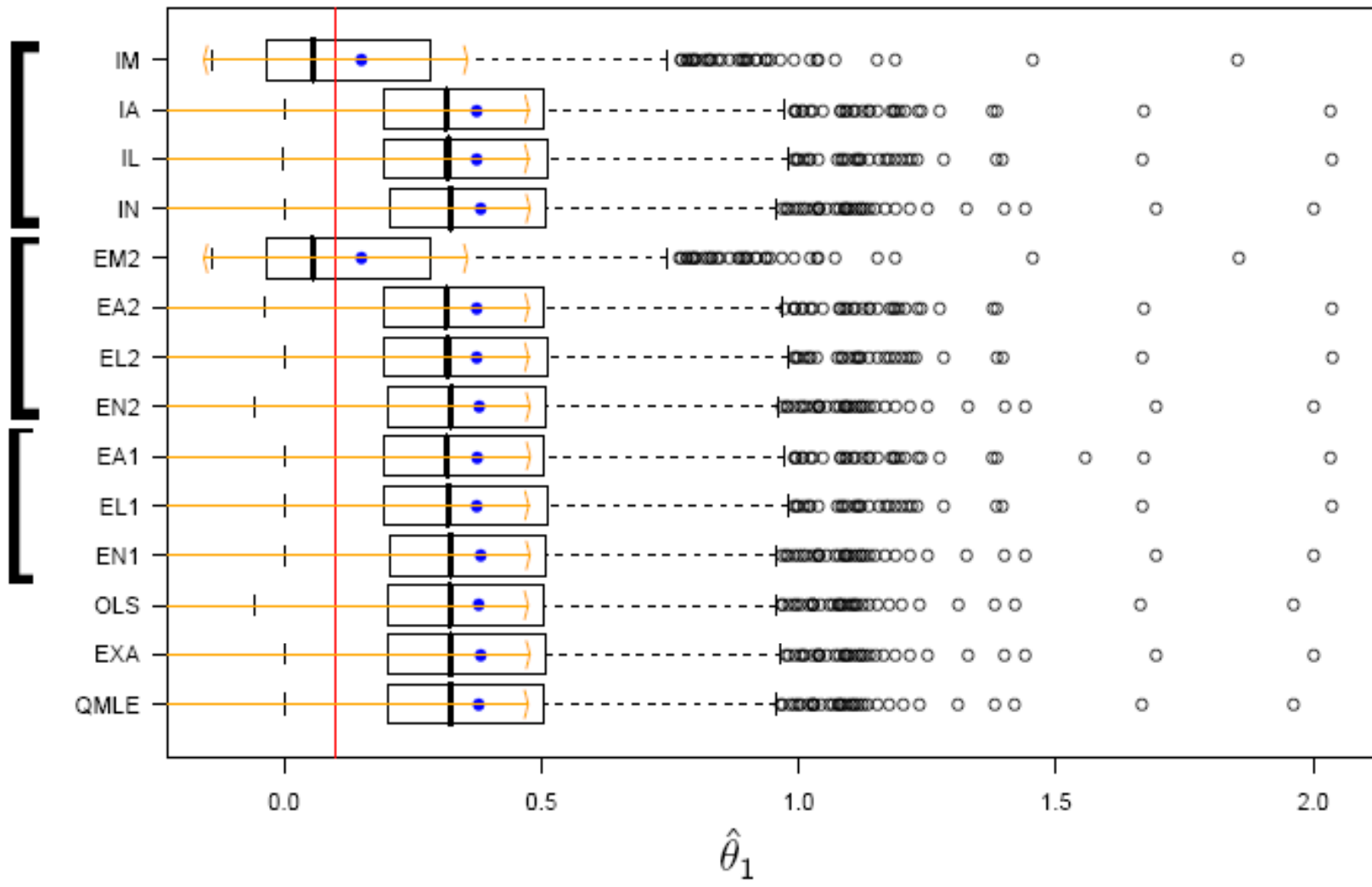


Note:  $S = 20$  for simulation-based estimates

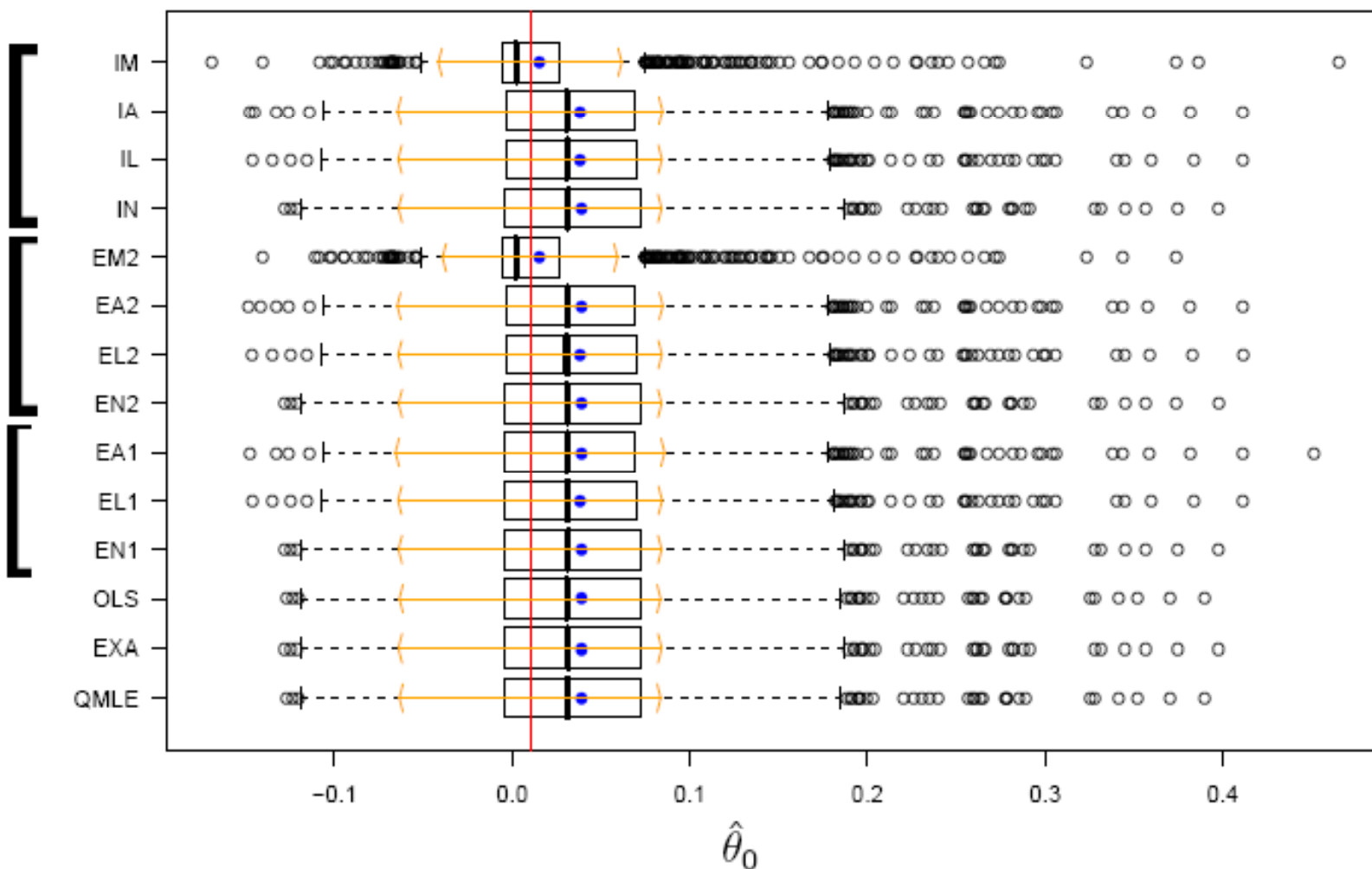
# Impact of simulation Size $S$ on Densities of $\theta_1$ Estimates



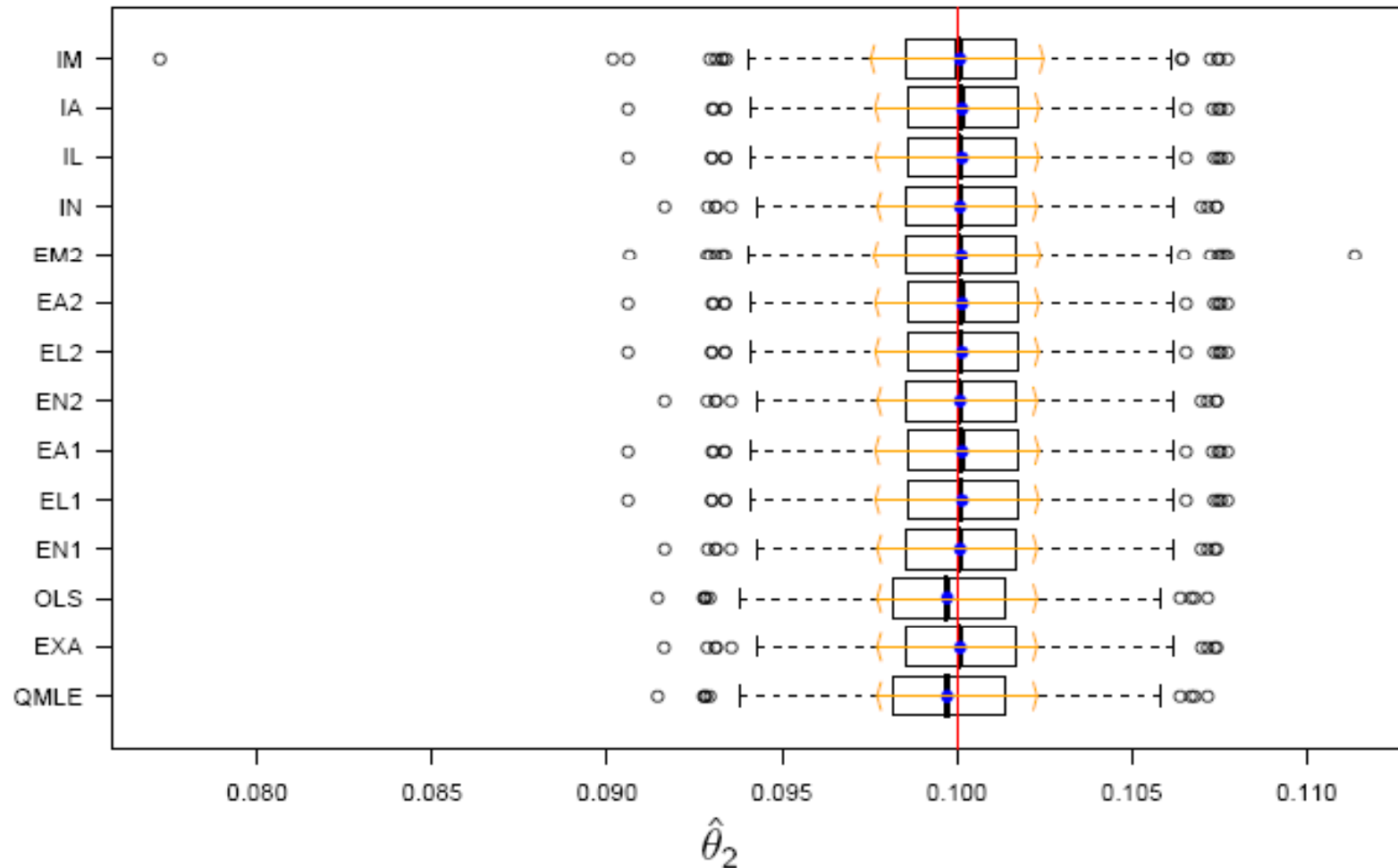
# Distribution of $\hat{\theta}_1$ : $\theta = (0.01, 0.1, 0.1)$ , $n=1000$ , $\Delta = 1/52$



# Distribution of $\hat{\theta}_0$ : $\theta = (0.01, 0.1, 0.1)$ , $n=1000$ , $\Delta = 1/52$



Distribution of  $\hat{\theta}_2$ :  $\theta = (0.01, 0.1, 0.1)$ ,  $n=1000$ ,  $\Delta = 1/52$





Rejection frequencies of likelihood ratio type tests  
at 5% nominal level of significance for 1000 Monte Carlo simulations.

<i>EN1</i>	<i>EL1</i>	<i>EA1</i>	<i>EN2</i>	<i>EL2</i>	<i>EA2</i>	<i>EM2</i>	<i>IN</i>	<i>IL</i>	<i>IA</i>	<i>IM</i>
True $\theta = (0.01, 0.1, 0.1)$ , $n = 1000$ , $\Delta = 1/52$ .										
Joint Test										
0.767	0.760	0.728	0.129	0.114	0.116	0.419	0.121	0.112	0.113	0.415
Simple Test of $\theta_1 = 0.1$										
0.698	0.640	0.593	0.197	0.169	0.175	0.130	0.210	0.176	0.183	0.149

# Research in Progress

- Fuleky, P., and Zivot, E. (2010). Further Evidence on Simulation Inference for Near Unit Root Processes with Implications for Term Structure Estimation. Manuscript in preparation.
- Fuleky, P., and Zivot, E. (2010). Indirect Inference Based on the Score. Manuscript in preparation.
- Fuleky, P., and Zivot, E. (2010). indirectInference: R package for indirect inference.

# References

- Duffee, G. and Stanton, R. (2008). Evidence on Simulation Inference for Near Unit-Root Processes with Implications for Term Structure Estimation. *Journal of Financial Econometrics*, 6(1):108.
- Gallant, A. and Tauchen, G. (1996). Which Moments to Match? *Econometric Theory*, 12(4):657-81.
- Lo, A. (1988). Maximum Likelihood Estimation of Generalized Ito Processes with Discretely Sampled Data. *Econometric Theory*, 4(2):231-247.
- Gouriéroux, C. and Monfort, A. (1996). *Simulation-Based Econometric Methods*. Oxford University Press, USA.

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- Gourieroux, C., Monfort, A., and Renault, E. (1993). Indirect Inference. *Journal of Applied Econometrics*, 8:S85-S118.
- Phillips, P. and Yu, J. (2009). Maximum Likelihood and Gaussian Estimation of Continuous Time Models in Finance. *Handbook of Financial Time Series*.
- Smith Jr, A. (1993). Estimating Nonlinear Time-Series Models Using Simulated Vector Autoregressions. *Journal of Applied Econometrics*, 8:S63-S84.